

1 **Predicting the impact of climate change on severe winter haze pollution events in**
2 **Beijing using extreme value theory**

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8 **Key Points:**

- 9 • Likelihood of extreme winter haze pollution in Beijing can be fit to meridional wind and
10 relative humidity using a point process model
- 11 • Application of the model to climate projections shows that the frequency of extreme haze
12 events is most likely to decrease by midcentury
- 13 • Disagreement with previous work projecting an increase in haze frequency is explained
14 by omission of RH from their models
15

16 Abstract

17 We use extreme value theory to develop point process (PP) statistical models relating the
18 probability of extreme winter pollution events in Beijing (winter haze) to local meteorological
19 variables. The models are trained with the 2009-2017 record of fine particulate matter
20 concentrations (PM_{2.5}) from the US embassy. We find that 850 hPa meridional wind velocity
21 (V850) and relative humidity (RH) successfully predict days when daily mean PM_{2.5} will exceed
22 300 $\mu\text{g m}^{-3}$ as well as higher thresholds. We apply the PP models to mid-21st century climate
23 model projections from the CMIP5 ensemble under two forcing scenarios (RCP8.5 and RCP4.5).
24 We find that the frequency of haze events is most likely to decrease because of climate change,
25 driven mainly by a decrease in RH. Our results contrast with previous studies that projected
26 increases in haze frequency but did not include RH as a meteorological predictor.

27 1 Introduction

28 Beijing in winter experiences very severe air pollution events, commonly referred to as ‘winter
29 haze.’ Concentrations of fine particulate matter (PM_{2.5}) can exceed 300 $\mu\text{g m}^{-3}$ on a 24-h average
30 basis. Efforts are underway to reduce emissions, but some studies have suggested that climate
31 change may partly offset the air quality gains by favoring the meteorological conditions leading
32 to winter haze [Cai *et. al.*, 2017; Zou *et. al.* 2017]. Here we use extreme value theory (EVT) to
33 investigate this climate penalty as it applies to severe pollution events.

34 Cai *et. al.* [2017] found that a meteorological index based on meridional wind velocity, zonal
35 wind gradient, and vertical temperature gradient was strongly correlated with winter haze days
36 exceeding 150 $\mu\text{g m}^{-3}$ in Beijing. They then used climate simulations from the IPCC CMIP5
37 multimodel ensemble driven by the RCP8.5 (business-as-usual) emissions scenario to project a
38 significant increase in this index over the 21st century, resulting in a 50% increase in haze event
39 frequency. Zou *et. al.* [2017] linked Arctic sea ice loss in the previous autumn and heavy
40 snowfall over Eurasia to poor ventilation conditions over the East China Plain, again implying
41 that 21st-century climate change would worsen winter haze. On the other hand, Horton *et. al.*
42 [2014] using the same CMIP5 RCP8.5 archive find no change between 1986-2005 and 2046-
43 2065 in the frequency of stagnation events over Beijing.

44 EVT offers a statistical framework to predict rare occurrences [Coles, 2001]. Here we model the
45 relationship of extreme haze days in Beijing with meteorological variables using an EVT
46 nonstationary Poisson point process (PP) model, which simultaneously fits the rate of PM_{2.5}
47 exceedances above a certain threshold and the probability density function (PDF) of PM_{2.5} above
48 that threshold. EVT has been applied to extreme pollution episodes in Switzerland, the US, and
49 South America [Rieder *et. al.*, 2010; Shen *et. al.*, 2016; Martins *et. al.*, 2017], but has not been
50 applied to Beijing haze to our knowledge.

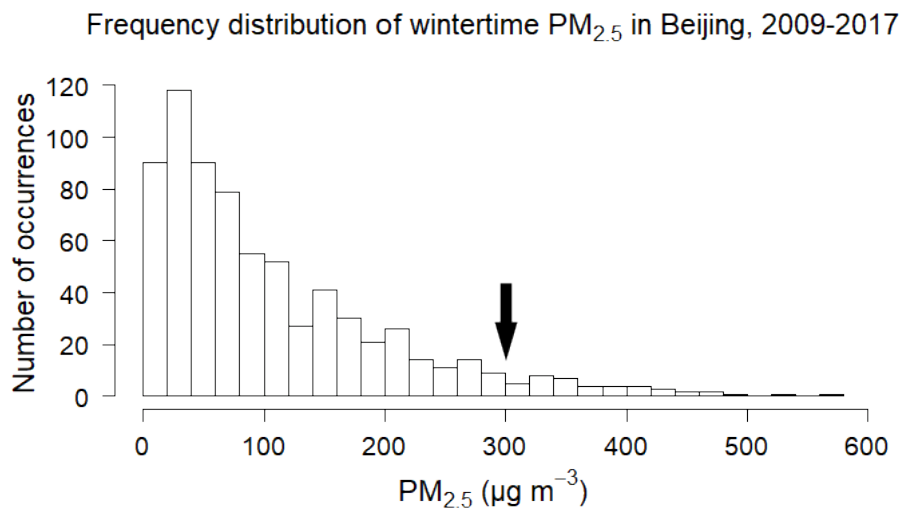
51 2 Data and Methods

52 We use 2009-2017 wintertime (DJF) time series of 24-h average PM_{2.5} in Beijing, starting in
53 December 2009 and ending in February 2017, measured at the US embassy and available from
54 (<http://www.stateair.net/web/historical/1/1.html>). Previous studies have shown that this dataset is
55 representative of PM_{2.5} in Beijing [San Martini *et. al.*, 2015; Mukherjee and Toohey, 2016]. Cai

56 *et al.* (2017) used it for their analysis of the relationship between Beijing haze events and
 57 meteorological variables.

58 We use the National Center for Environmental Prediction and National Center for Atmospheric
 59 Research (NCEP/NCAR) reanalysis at 2.5° by 2.5° resolution [Kalnay *et al.*, 1996] to construct
 60 the meteorological covariates of wintertime Beijing haze previously identified by Cai *et al.*
 61 [2017] in their Haze Weather Index. These include (1) 850 hPa meridional wind velocity (V850),
 62 averaged spatially in the area around and to the southeast of Beijing (30° – 47.5° N, 115° – 130° E),
 63 (2) the difference in the 500hPa zonal wind (δU_{500}) north of Beijing (42.5° – 52.5° N, 110° –
 64 137.5° E) vs. south of Beijing (27.5° – 37.5° N, 110° – 137.5° E), and (3) the vertical temperature
 65 difference ($\delta T_{850-250}$) between the lower troposphere at 850 hPa (32.5° – 45° N, 112.5° – 132.5° E)
 66 and the upper troposphere at 250 hPa (37.5° – 45° N, 122.5° – 137.5° E). We also consider (4) the
 67 relative humidity (RH) at Beijing International airport obtained from the NOAA Integrated
 68 Surface Database (<https://www.ncdc.noaa.gov/isd/data-access>), since several previous studies
 69 have reported a positive correlation of Beijing haze with RH [Wang *et al.*, 2014; Zheng *et al.*,
 70 2015; Lueng *et al.*, 2018; Shen *et al.*, 2018]. The Beijing airport data is more reliable for local
 71 RH than the NCEP/NCAR reanalysis [Shen *et al.*, 2018]. All meteorological variables are used
 72 as 24-h averages matched to the $PM_{2.5}$ observations.

73 Figure 1 shows the PDF of wintertime 24-h average $PM_{2.5}$ in Beijing for 2009-2017. We fit the
 74 high tail of the PDF to a Poisson point process (PP) model. EVT holds that if we consider a
 75 sufficiently large number of independent, identically distributed measurements of a random
 76 variable X and restrict our attention to the codomain $[u, \infty)$ for a sufficiently large threshold u ,
 77 the PDF of $X \in [u, \infty)$ converges to a PP model regardless of the original underlying distribution
 78 of X [Coles 2001]. Our PP model simulates the Poisson process limit of 24-hour average $PM_{2.5}$ in
 79 Beijing above a $u = 300 \mu\text{g m}^{-3}$ threshold, representing the 95th percentile of the data (Figure 1).



80

81 **Figure 1.** Frequency distribution of daily wintertime (DJF) 24-hour average $PM_{2.5}$ in
 82 Beijing for 2009-2017, starting in December 2009 and ending in February 2017. Data are
 83 from the US embassy site. The arrow indicates the $300 \mu\text{g m}^{-3}$ threshold used in our

84 definition of extreme events and representing the 95th percentile of the frequency
85 distribution.

86 We can fit the PP model to different meteorological predictors. Consider a fit PP(V850, RH) to
87 V850 (v) and RH (r), the two meteorological variables identified in *Shen et. al.* [2018] as the
88 strongest predictors of Beijing haze. Let y denote 24-h average PM_{2.5}. The probability that y will
89 exceed a threshold u given (v, r) is modeled with the marginal distribution

$$90 \quad P(y > u | v, r) = \frac{1}{n_a} \left(1 + \xi \left(\frac{u - \mu_{v,r}}{\phi_v} \right) \right)^{-\frac{1}{\xi}} \quad (1)$$

91 with

$$92 \quad \mu_{v,r} = av + br + c \quad (2)$$

$$93 \quad \phi_v = e^{dv+f} \quad (3)$$

94 Here $\mu_{v,r}$ is the location parameter written as a linear function of v and r , ϕ_v (> 0) is the scale
95 parameter written as an exponential function of v , and ξ is the shape parameter which describes
96 the shape of the distribution rather than shifting it (as $\mu_{v,r}$ does) or stretching it (as ϕ_v does). n_a
97 represents the number of observations in a year (90, excluding leap days). a, b, c, d, f, ξ are
98 model parameters to be optimized. This optimization is done by maximizing the likelihood
99 function L [Coles, 2001]:

$$100 \quad L(\mu_{v,r}, \phi_v, \xi) = \underbrace{\exp \left(-\frac{1}{n_a} \sum_{t=1}^n \left(1 + \frac{\xi(u - \mu_{v,r})}{\phi_v} \right)^{-\frac{1}{\xi}} \right)}_{(a)} \underbrace{\prod_{t=1}^n \left(\frac{1}{\phi_v} \left(1 + \frac{\xi(y_t - \mu_{v,r})}{\phi_v} \right)^{-\frac{1}{\xi}-1} \right)^{I(y_t > u)}}_{(b)} \quad (4)$$

101 where the summation is done over all observations (y_t, v_t, r_t) at times $t \in [1, n]$ from 2009 to
102 2017. Here $I(y_t > u)$ is the indicator function equal to 1 if $y_t > u$ and 0 otherwise. L is a
103 product of two factors: factor (b) represents exceedances ($y_t > u$) explicitly as a product of
104 independent generalized Pareto densities, while factor (a) incorporates information from all
105 observations, even those that are not extreme. L is maximized when $\mu_{v,r}$, ϕ_v , and ξ best explain
106 the data. The optimization is performed with extRemes, an EVT package in the statistical
107 software R [Gilleland and Katz, 2011]. The forms of $\mu_{v,r}$ and ϕ_v in equations (2) and (3) were
108 chosen after testing a range of possible forms (such as $\phi = e^{dv+fr+g}$) to maximize the value of
109 L subject to the Akaike Information Criterion (AIC). The AIC penalizes models with a large
110 number of optimized parameters p and rewards those with a high maximum likelihood value L :

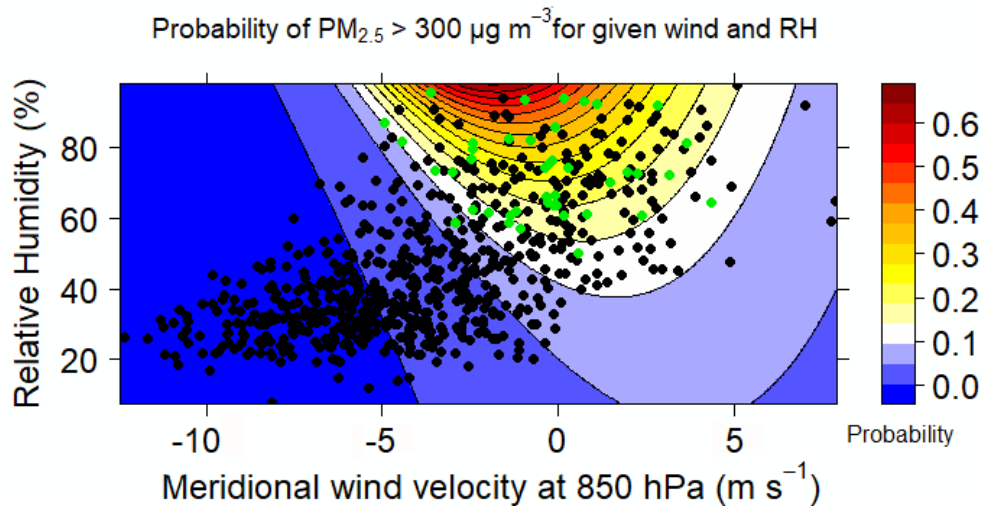
$$111 \quad \text{AIC} = 2p - 2 \ln L \quad (5)$$

112 The model with the lowest AIC is preferred [Aikake, 1974].

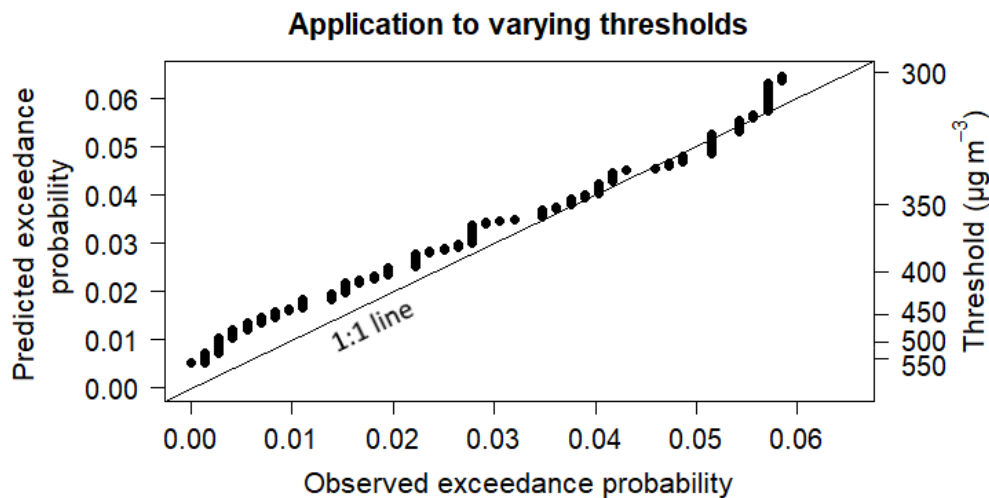
113 Results and Discussion

114 **Figure 2 (top panel)** shows the model probabilities for $\text{PM}_{2.5}$ to exceed the $u = 300 \mu\text{g m}^{-3}$
 115 threshold as a function of V850 and RH. The model matches the observed occurrences.
 116 Maximum probabilities are for near-zero wind (stagnant conditions) and high RH (conducive to
 117 PM formation; Wang *et al.* [2014]).

118 An important property of the PP model is threshold invariance. Although we fit the model with
 119 $u = 300 \mu\text{g m}^{-3}$, EVT predicts that the same parameters $\mu_{v,r}$, ϕ_v , and ξ should yield accurate
 120 estimates for any value $u > 300 \mu\text{g m}^{-3}$ in equation (1) [Coles, 2001]. This is verified in the
 121 bottom panel of Figure 2, and the implication is that the model describes not only the probability
 122 of exceeding $300 \mu\text{g m}^{-3}$ but also the PDF above that threshold.



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125 **Figure 2.** Probabilities of exceedance of extreme daily mean $\text{PM}_{2.5}$ thresholds in
 126 wintertime (DJF) Beijing haze, as described by an extreme value point process model
 127 PP(V850, RH) conditioned on 850 hPa meridional wind velocity (V850, positive
 128 northward) and relative humidity (RH). Colored contours in the top panel shows the
 129 model probability of $\text{PM}_{2.5}$ exceeding $300 \mu\text{g m}^{-3}$ as a function of V850 and RH.
 130 Symbols show the 2009-2017 observations, with observed exceedances of the $300 \mu\text{g m}^{-3}$

131 threshold indicated in green. Each symbol represents average V850 and RH values on
 132 one day. The bottom panel shows the ability of the same model to reproduce the observed
 133 probability of occurrence of higher thresholds. The identity line is plotted for reference.

134 We also attempted to fit our EVT model to all four meteorological covariates identified in
 135 **Section 2** for Beijing haze: V850, RH, δU_{500} , and $\delta T_{850-250}$. However, this was unsuccessful
 136 because of strong correlations between the variables. To get around this problem, we performed
 137 a principal component analysis (PCA) of the four variables and found that the first two principal
 138 components account for 90% of the variance. Thus we fit the PP model to just these two
 139 components. The first principal component gives an almost-equal weighting to δU_{500} , V850,
 140 and $\delta T_{850-250}$, with RH counting for slightly less; the second principal component is dominated by
 141 RH. Best performance for the EVT model takes the location parameter as a linear combination of
 142 the first two principal components and the scale parameter as an exponential function of the first
 143 principal component. The resulting probability heatmap and threshold exceedance plot for this
 144 PP(V850, RH, δU_{500} , $\delta T_{850-250}$) model are given in **Figure S1**.

145 For direct comparison with *Cai et. al.*, [2017], we fit yet a third model using the three covariates
 146 identified in that paper (V850, δU_{500} , and $\delta T_{850-250}$). A PCA shows that the first principal
 147 component accounts for 84% of the variance. We fit our PP(V850, δU_{500} , $\delta T_{850-250}$) model with
 148 just that vector as a covariate. The location parameter is again a linear function and the scale
 149 parameter is an exponential.

150 By calculating and comparing the AIC for these three candidate PP models (AIC_i with $i \in$
 151 $\{1,2,3\}$), we can determine which model i minimizes information loss. We do this by comparing
 152 Aikake weights w_i , given by

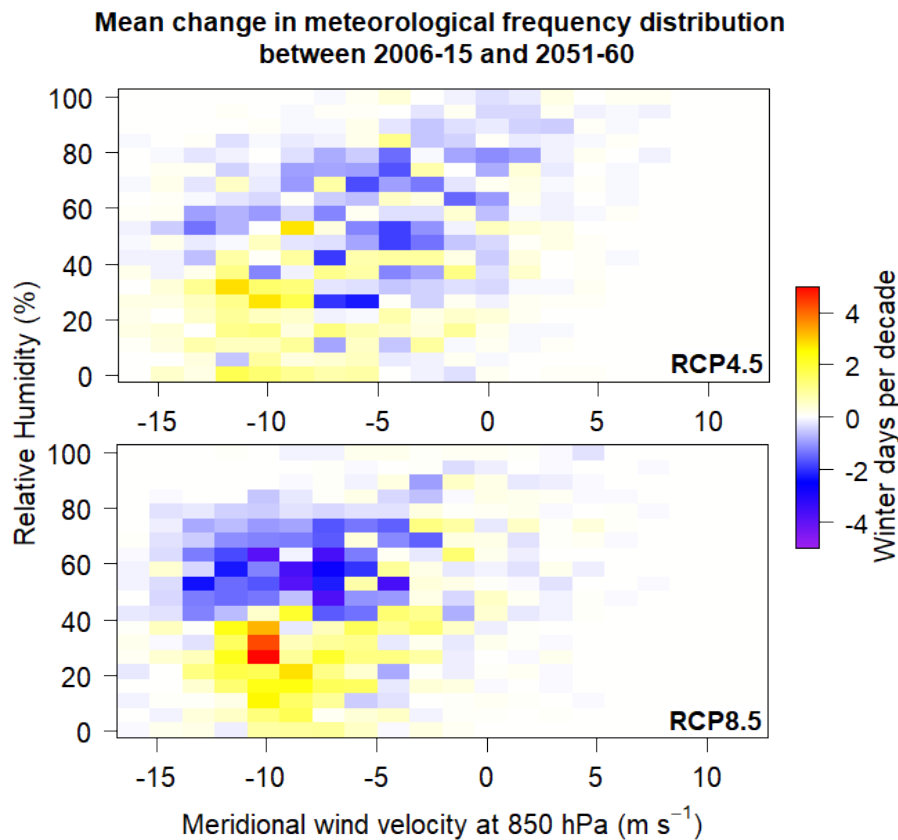
$$153 \quad w_i = \frac{\exp\left(\frac{1}{2}(AIC_{\min} - AIC_i)\right)}{\sum_{j=1}^3 \exp\left(\frac{1}{2}(AIC_{\min} - AIC_j)\right)} \quad (6)$$

154 where $AIC_{\min} = \min(AIC_i)$. The Aikake weights represent the relative likelihood of a model
 155 given the data, and for our three models are 0.87, 0.13, and 0.00 respectively for PP(V850, RH),
 156 PP(V850, RH, δU_{500} , $\delta T_{850-250}$), and PP(V850, δU_{500} , $\delta T_{850-250}$). This provides strong evidence
 157 for preferring the PP(V850, RH) model, a conclusion supported by previous work pointing to
 158 V850 and RH as the most powerful predictors of Beijing haze [*Zheng et. al.*, 2015; *Shen et. al.*,
 159 2018].

160 We can now project the effect of 2006-2060 climate change on the frequency of extreme Beijing
 161 haze events by applying our PP(V850, RH) model to daily meteorological projections from the
 162 ensemble of 18 global climate models participating in the IPCC Coupled Model Intercomparison
 163 Project (CMIP5, see **Table S1** for details) that report daily V850 and RH data. To obtain a
 164 timeseries for future RH, we average spatially over the same region used to calculate V850 (30°–
 165 47.5° N, 115°–130° E). Present-day RH statistics in these CMIP5 models often differ
 166 significantly from the Beijing airport data used to fit our PP model. For consistency in
 167 application of the PP model, we adjust the 2006-2015 meteorological variables for each CMIP5

168 model to match the observed means and standard deviations used in the PP model, and we apply
 169 the same adjustment to the CMIP5 model values through to 2060.

170 **Figure 3** shows the two-dimensional frequency distribution of mean projected changes of DJF
 171 V850 and RH for the ensemble of models under the two scenarios of climate forcing and on the
 172 same scale as Figure 2. Here we take the change from the 2006-15 period to 2051-60 for each
 173 model and then average these changes over the models. Results are shown for the RCP4.5 and
 174 RCP8.5 radiative scenarios, where the RCP4.5 scenario describes a more moderate climate
 175 forcing than the RCP8.5 business-as-usual scenario. We use 18 models from the RCP8.5
 176 scenario and 16 from RCP4.5, reflecting the variables available.



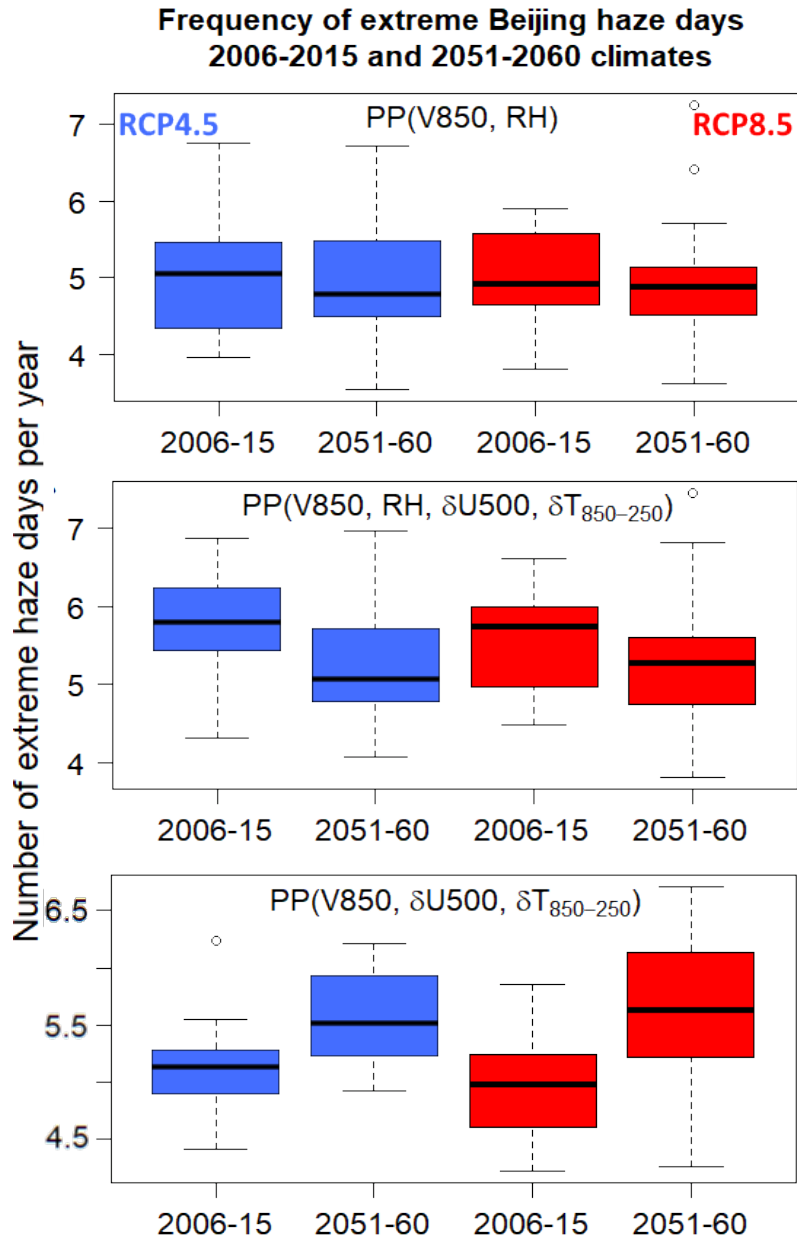
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178 **Figure 3.** Projected changes in the wintertime (DJF) meridional wind velocity at 850 hPa
 179 (V850) and relative humidity (RH) over Beijing from 2006-2015 to 2051-2060. Results
 180 are from the ensemble of CMIP5 climate models reporting daily V850 and RH, and for
 181 two climate forcing scenarios (RCP4.5 and RCP8.5). Values are changes in the number
 182 of winter days per decade for each (V850, RH) bin, computed for individual models and
 183 then averaged across models. The scale is the same as for the extreme haze probability
 184 map in Figure 2.

185 We find in Figure 3 a significant decrease in RH from 2006-2015 to 2015-2060 but no
 186 significant change in V850. These trends are broadly consistent with trends in eastern
 187 China as a whole [Shen *et al.*, 2018]. The decrease in RH under greenhouse warming may
 188 be expected from increased land-ocean temperature contrast and poleward expansion of

189 the Hadley cell [Byrne et al., 2013; Lau et al., 2015]. There is however substantial
190 variability between models including in the sign. For the RCP4.5 scenario 63% of the
191 CMIP5 models project a decrease in mean RH, and 75% project a decrease in the
192 frequency of days where RH is greater than 80%. For the RCP8.5 scenario 78% project a
193 decrease in mean RH, and 61% project a decrease in the frequency of RH greater than
194 80%.

195 The trends in Figure 3 computed for individual CMIP5 climate models can be combined
196 with the extreme haze probabilities in Figure 2 to project the effect of climate change on
197 extreme haze for the PP(V850, RH) model, which we have shown above to be most
198 reliable. Results are shown in Figure 4 (top panel) for the ensemble of CMIP5 models.
199 The model medians show a decrease in both the RCP4.5 and RCP8.5 scenarios. By
200 calculating a 2006-2060 linear regression of the annual number of haze days averaged
201 over the ensemble of models, we find the decrease in the RCP4.5 scenario is significant
202 at the 5% level and amounts to 0.58 ± 0.39 fewer haze days per year in 2060 than in
203 2006, representing a 10% decrease. The decrease in the RCP8.5 scenario is not
204 statistically significant. In the RCP4.5 scenario, 69% of the CMIP5 models show a
205 decrease in the frequency of haze days. In the RCP8.5 scenario, 56% of the models show
206 a decrease.



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Figure 4. Change in the number of extreme Beijing haze days per year for the 2051-2060 versus 2006-2015 climate as predicted from extreme value theory using climate projections from 18 CMIP5 models. Extreme haze days are defined by 24-hour average $PM_{2.5} > 300 \mu g m^{-3}$ and are assumed to occur in winter (DJF) only. The three panels show predictions from different Poisson point process (PP) models fit to different combinations of local meteorological variables including meridional wind velocity at 850 hPa (V850), relative humidity (RH), meridional gradient in 500 hPa zonal wind velocity (δU_{500}), and vertical temperature difference between 850 and 250 hPa ($\delta T_{850-250}$). The boxplots show statistics for the ensemble of CMIP5 models: medians, 25th and 75th percentiles, and ranges. All changes with the exception of RCP8.5 in the topmost panel are statistically significant at the 95% level (see text)

219 The middle and bottom panels of Figure 4 show results from the alternative PP models.
 220 The PP(V850, RH, δU_{500} , $\delta T_{850-250}$) model shows significant decreases in the frequency
 221 of haze days by 2060 for both the RCP4.5 and RCP8.5 scenarios, by -0.87 ± 0.25 days
 222 (10%) and -0.54 ± 0.20 days (10%) respectively for the two scenarios. By contrast, the
 223 PP(V850, δU_{500} , $\delta T_{850-250}$) model shows a significant increase for both scenarios, by
 224 0.87 ± 0.27 days (20%) for RCP8.5. This is consistent with *Cai et al.* [2017], who found
 225 from the same three variables in the RCP8.5 scenario a 50% increase in the frequency of
 226 haze days $> 150 \mu\text{g m}^{-3}$ between the 1950-1999 and 2050-2099 climates (as compared to
 227 2006-2015 vs. 2051-2060 in our work). As pointed out above, a major flaw of this
 228 PP(V850, δU_{500} , $\delta T_{850-250}$) model is that it does not include RH as a predictor variable.

229 We repeated the statistical comparison of haze frequency between the 2006-2015 and 2051-2060
 230 climates for two higher thresholds, 350 and 400 $\mu\text{g m}^{-3}$, to check the sensitivity of our results to
 231 threshold selection. For all three PP models, the significance and sign of the changes in the
 232 frequency of haze days was consistent at these higher thresholds.

233 In summary, we find from extreme value theory that the probability of severe wintertime haze
 234 events in Beijing (24-h average $\text{PM}_{2.5} > 300 \mu\text{g m}^{-3}$) can be successfully represented by a point
 235 process (PP) model with meridional wind velocity (V850) and relative humidity (RH) as the two
 236 meteorological predictors. The probability of extreme events is highest under near-zero V850
 237 and high RH, as would be expected from stagnation and conditions conducive to aqueous aerosol
 238 formation. Alternative PP models including as additional meteorological variables the gradient in
 239 500 hPa zonal wind (δU_{500}) and the vertical temperature gradient ($\delta T_{850-250}$) are not as
 240 successful. RH is a particularly important predictor. Application of the PP(V850, RH) model to
 241 2006-2060 CMIP5 simulations of climate change following the RCP4.5 scenario reveals a 10%
 242 decrease in the probability of extreme events given constant emissions; the decrease is driven
 243 mainly by decreasing RH. The RCP8.5 scenario shows no significant trend. A PP(V850, RH,
 244 δU_{500} , $\delta T_{850-250}$) model shows significant decreases in haze frequency in both scenarios. In
 245 contrast, a PP(V850, δU_{500} , $\delta T_{850-250}$) model not including RH as predictor variable shows an
 246 increase in haze days for both the RCP4.5 and RCP8.5 scenarios, consistent with the results of
 247 *Cai et al.* [2017]. However, we find in evaluating model success that RH needs to be included as
 248 a predictor variable since extreme haze events are strongly associated with high RH. Increasing
 249 greenhouse forcing may be expected to cause a decrease of RH over China due to increased land-
 250 ocean temperature contrast and the poleward expansion of the Hadley circulation. We conclude
 251 that 21st-century climate change should not increase the frequency of extreme haze events in
 252 Beijing and is more likely to decrease it.

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 258 Programme's Working Group on Coupled Modelling, which is responsible for CMIP, and we
 259 thank the climate modeling groups listed in **Table S1**. For CMIP the U.S. Department of
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262 Organization for Earth System Science Portals.

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